

Errata
Analysis of Panel Data
2nd Edition

1. p. 39, eq. (3.3.20):

$$\frac{\partial \log L}{\partial \tilde{\beta}} = \frac{1}{\sigma_u^2} \left[\sum_{i=1}^N (y_i - e\mu - X_i \tilde{\beta})' Q X_i + \frac{T\sigma_u^2}{\sigma_u^2 + T\sigma_\alpha^2} \sum_{i=1}^N (\bar{y}_i - \mu - \bar{x}_i' \tilde{\beta}) \bar{x}_i' \right] = \underline{0}.$$

2. p. 87, line 6: respectively, and ΔX_i is the $(T-1) \times K_1$ matrix of ...

3. p. 109, first line after eq. (4.7.27):
 where $\tilde{r}_i = (\Delta y_i - e\delta)$,

4. p. 189, the first line after eq. (7.2.2): Then the expected value of y is the probability ... (take out the subscript i from y_i)

5. p. 189, eq. (7.2.6):

$$\begin{aligned} Pr(y = 1 | \tilde{x}) &= \int_{-\tilde{\beta}'\tilde{x}}^{\infty} \frac{\exp(\nu)}{(1 + \exp(\nu))^2} d\nu \\ &= \frac{\exp(\tilde{\beta}'\tilde{x})}{1 + \exp(\tilde{\beta}'\tilde{x})}. \end{aligned}$$

6. p. 189, first line after eq. (7.2.6):

Let $F(\tilde{\beta}'\tilde{x}) = E(y | \tilde{x})$. (take out the subscript i from y_i)

7. p. 201, line 16: replace the "=" by "≈".

8. p. 212, line 6 to the end of this paragraph (1.7): ..., then the conditional probability changes according to $y_{i1} = 0$ or 1.

9. p. 229, line 1: in the limit. There could be multiple roots that satisfy (8.1.12) or (8.1.13). For instance, $\hat{\beta} = \underline{0}$ is one such root. To ensure uniqueness, Powell (1986) proposes the following symmetrically trimmed least squares estimator...

10. p. 229, third line after eq. (8.1.15): The motivation for $S_N(\tilde{\beta})$ is that observations greater than $2\tilde{\beta}'\tilde{x}$ if $\tilde{\beta}'\tilde{x} > 0$ will have partial derivatives equal to $(\tilde{\beta}'\tilde{x})\tilde{x}$ and ...

11. p. 232, last line: ... by a weighted least squares (or instrumental variables) reregression.
12. p. 250, line 26: ... $A_2 = \{(y_{i1}^*, y_{i2}^*) : y_{i1}^* \leq 0, y_{i2}^* > \Delta x_{i2}^* \beta\}$
13. p. 251, line 3, eq. (8.4.21): $\times (y_{i2} - \max(0, \Delta x_{i2}^* \beta)) + 1 \{(y_{i1}, y_{i2}) \in B_1\} (\Delta y_i - \Delta x_{i2}^* \beta)$
14. p. 264, first line below eq. (8.6.12): when $\Gamma = \frac{\partial}{\partial \theta} E[m(\theta)]$,
15. p. 264, line 5 from the last paragraph: 8.6 and 8.7 approaches zero. ...